Cross-validation Confidence Intervals for Test Error

Alexandre Bayle

Harvard University



Joint work with **Pierre Bayle** (Princeton University), **Lucas Janson** (Harvard University), and **Lester Mackey** (Microsoft Research)

November 5, 2020

k-fold cross-validation



- Unbiased
- Lower variance than a single train-test split
- Complex dependence structure

High-stakes applications

Prediction of cancer outcome with microarrays: a multiple random validation strategy

Lavor 2005; 265-488-42 Stefon Michiels, Serge Koscielry, Catherine Hill

See Comment page 454 Bestatistics and Epidemology Summary

Unt 5 Mobile 105, 5 Mobile Background General studies of microarray gene-expression profiling have been undertaken to predict cancer no. CHINE, functional outcome. Knowledge of this gene-expression profile or molecular signature should improve treatment of patients by commission profile or molecular signature should improve treatment of patients by newsilk/integrate and allowing treatment to be tailored to the severity of the disease. We reanalysed data from the seven largest published incruct outaw kossy, studies that have attempted to predict prognosis of cancer patients on the basis of DNA microarray analysis. Mileful France

Convegendences: Methods The standard strategy is to identify a molecular signature (ie, the subset of genes most differentially to sepressed in patients with different outcomes) in a training set of patients and to estimate the proportion of expressed in patients with different outcomes) in a training set of patients and to estimate the properties of an independent validation set of patients. We expanded this strategy from the strategy for testing the strategy for Demode, staty/mp.r. (based on unique training and validation sets) by using multiple random sets, to study the stability of the molecular signature and the proportion of misclassifications.

Findings The list of genes identified as predictors of prognosis was highly unstable; molecular signatures strongly depended on the selection of patients in the training sets. For all hut one study, the proportion misclassified decreased as the number of patients in the training set increased. Because of inadequate validation, our chosen studies published overoptimistic results compared with those from our own analyses. Five of the seven studies did not classify patients better than chance.

Interpretation The prognostic value of published microarray results in cancer studies should be considered with caution. We advocate the use of validation by repeated random sampling.

Introduction

guidelines (Minimum Information About a Microarray The expression of several thousand genes can be studied Experiment'). This approach offers an opportunity to simultaneously by use of DNA microarrays. These propose alternative analyses of these data. We have taken microarrays have been used in many specialties of advantage of this opportunity to analyse different datasets medicine. In oncology, their use can identify genes with from rublished studies of gene expression as a predictor

High-stakes applications

Articles

Prediction of cancer outcome with microarrays: a multiple random validation strategy

Lanon 2005; 355: 480-92 Stefon Michiels, Serge Koscielity, Catherine Hill

Statistical Applications in Genetics and Molecular Biology

Volume 7, Issue 1	2008	Article 8
-------------------	------	-----------

Calculating Confidence Intervals for Prediction Error in Microarray Classification Using Resampling

Wenyu Jiang, Concordia University Sudhir Varma, Genomics and Bioinformatics Group, Laboratory of Molecular Pharmacology, National Cancer Institute Richard Simon, Biometric Research Branch, Division of Cancer Treatment and Diagnosis, National Cancer Institute array gene-expression profiling have been undertaken to predict cancer sion profile or molecular signature should improve treatment of patients by severity of the disease. We reanalysed data from the seven largest published rognosis of cancer patients on the basis of DNA microarray analysis.

lentify a molecular signature (ie, the subset of genes most differentially utcomes) in a training set of patients and so estimate the proportion of on an independent validation set of patients. We expanded this strategy tion sets) by using multiple random sets, to study the stability of the 1 of misclassifications.

peedictors of prognosis was highly unstable; indecular signatures strengly in the training sets. For all but one study, the proportion mixclassified a the training set increased. Because of inadequex validation, our chosen ts compared with those from our own analyses. Five of the seven studies unce.

published microarray results in cancer studies should be considered with ion by repeated random sampling.

s can be studied recurrently. This approach offers an opportunity to recurrently. These propose aberrative analyses of these data. We have taken y specialities of advantage of this opportunity to analyse different datasets midif genes with from published studies of gene expression as a predictor

High-stakes applications

Prediction of cancer outcome with microarrays: a multiple random validation strategy

Lawor 2005; 265-488-62 Stefon Michiels, Serge Koscielny, Catherine Hill

Statistical Applications in Genetics and Molecular Biology

Volume 7, Issue 1	2008	Article 8
-------------------	------	-----------

Calculating Confidence Intervals for Prediction Error in Microarray Classification Using Resampling

Wenvu Jiang, Concordia University Sudhir Varma, Genomics and Bioinformatics Group. Laboratory of Molecular Pharmacology, National Cancer Institute Richard Simon, Biometric Research Branch, Division of Cancer Treatment and Diagnosis. National Cancer Institute array gene-expression profiling ssion profile or molecular signatu severity of the disease. We reanab rognosis of cancer patients on the

lentify a molecular signature (in stcomest in a training set of pa on an independent validation se tion sets) by using multiple ra-1 of misclassifications.

predictors of prognosis was highl in the training sets. For all but the training set increased. Beca ts compared with those from ou

published microarray results in ion by repeated random samplin

suidelines (A is can be studied Emeriment? roarrays. These propose altern y specialties of advantage of t ntify senes with from rublish

NIH Public Access Author Manuscript

Published in final edited form as

Loncet Reprir Med. 2015 January : 3(1): 42-52. doi:10.1016/82213-2600(14)70239-5.

Mortality prediction in the ICU: can we do better? Results from the Super ICU Learner Algorithm (SICULA) project, a populationbased study

Romain Pirracchio, MD^{13,0} Maya L. Petersen, MD¹ Marco Carone, PhD² Matthieu Resche Rigon, MD⁵, Prof. Sylvie Chevret, MD⁵, and Prof. Mark J. van der LAAN, PhD

"Division of Biostatistics, School of Public Health, University of California, Berkeley, Berkeley, USA ⁸Service de Biostatistique et Information Médicale, Unité INSERM 1153, Equipe ECSTRA,

Service d'Anesthésie-Réanimation, Hópital Européen Georges Pompidou, Université Paris 5 Descartes, Sorbonne Paris Cité, Paris, France

*Department of Biostatistics, School of Public Health, University of Washington, Seattle, USA

Abstract

Background-Improved mortality prediction for patients in intensive care units (ICU) remains an important challenge. Many severity scores have been proposed but validation studies have concluded that they are not adequately calibrated. Many flexible algorithms are available, yet none of these individually outperform all others recordless of context. In contrast, the Super Learner (SL), an ensemble machine learning technique that leverages on multiple learning absorithms to obtain better prediction performance, has been shown to perform at least as well as the optimal member of its library. It might provide an ideal opportunity to construct a novel severity score with an improved performance profile. The aim of the present study was to provide a new

Is algorithm A actually better than algorithm B?

• $(Z_i)_{i\geq 1} = (X_i,Y_i)_{i\geq 1}$, X_i vector of covariates and Y_i target variable

- $(Z_i)_{i\geq 1} = (X_i,Y_i)_{i\geq 1}$, X_i vector of covariates and Y_i target variable
- For any vector B of indices in $\{1, \ldots, n\}$, let Z_B denote the subvector of $Z_{1:n}$ corresponding to ordered indices in B.

- $(Z_i)_{i\geq 1} = (X_i,Y_i)_{i\geq 1}$, X_i vector of covariates and Y_i target variable
- For any vector B of indices in $\{1, \ldots, n\}$, let Z_B denote the subvector of $Z_{1:n}$ corresponding to ordered indices in B.
- Consider a set of k train-validation splits $\{(B_j, B'_j)\}_{j=1}^k$ with validation indices $\{B'_j\}_{j=1}^k$ partitioning $\{1, \ldots, n\}$ into k folds.

- $(Z_i)_{i\geq 1} = (X_i, Y_i)_{i\geq 1}$, X_i vector of covariates and Y_i target variable
- For any vector B of indices in $\{1, \ldots, n\}$, let Z_B denote the subvector of $Z_{1:n}$ corresponding to ordered indices in B.
- Consider a set of k train-validation splits $\{(B_j, B'_j)\}_{j=1}^k$ with validation indices $\{B'_j\}_{j=1}^k$ partitioning $\{1, \ldots, n\}$ into k folds.
- $h_n(Z_i, Z_B)$: scalar loss function, evaluating the loss on the test point $Z_i = (X_i, Y_i)$ of the prediction rule learned on the training data Z_B . Examples of h_n :

$$h_n(Z_i, Z_B) = \begin{cases} (Y_i - \hat{f}(X_i; Z_B))^2 \\ \mathbb{1}[Y_i \neq \hat{f}(X_i; Z_B)] \end{cases}$$

k-fold cross-validation error and k-fold test error

Definition (*k*-fold cross-validation error)

$$\hat{R}_n \triangleq \frac{1}{n} \sum_{j=1}^k \sum_{i \in B'_j} h_n(Z_i, Z_{B_j})$$

- k either fixed or dependent on n
- The terms are not independent. What is the asymptotic behavior?

k-fold cross-validation error and k-fold test error

Definition (*k*-fold cross-validation error)

$$\hat{R}_n \triangleq \frac{1}{n} \sum_{j=1}^k \sum_{i \in B'_i} h_n(Z_i, Z_{B_j})$$

- k either fixed or dependent on n
- The terms are not independent. What is the asymptotic behavior?

Definition (*k*-fold test error)

$$R_n \triangleq \frac{1}{n} \sum_{j=1}^k \sum_{i \in B'_j} \mathbb{E}[h_n(Z_i, Z_{B_j}) \mid Z_{B_j}]$$

It is a standard inferential target and represents the average test error of the k prediction rules $\hat{f}(\cdot; Z_{B_j})$ for $j = 1, \ldots, k$.

k-fold cross-validation error and k-fold test error

Definition (*k*-fold cross-validation error)

$$\hat{R}_n \triangleq \frac{1}{n} \sum_{j=1}^k \sum_{i \in B'_i} h_n(Z_i, Z_{B_j})$$

- k either fixed or dependent on n
- The terms are not independent. What is the asymptotic behavior?

Definition (*k*-fold test error)

$$R_n \triangleq \frac{1}{n} \sum_{j=1}^k \sum_{i \in B'_j} \mathbb{E}[h_n(Z_i, Z_{B_j}) \mid Z_{B_j}]$$

It is a standard inferential target and represents the average test error of the k prediction rules $\hat{f}(\cdot; Z_{B_j})$ for $j = 1, \ldots, k$.

Goal

Central Limit Theorem on \hat{R}_n under mild assumptions

Stability

How much does the performance of a learned prediction rule change when one point in the training set is changed? Different kinds of stability, for example:

- Uniform stability
- Mean-square stability γ_{ms}
- Loss stability γ_{loss}

Note: $\gamma_{loss} \leq \gamma_{ms}$

[Bousquet and Elisseeff, 2002, Kale et al., 2011, Kumar et al., 2013, Celisse and Guedj, 2016, \dots]

Theorem

Suppose $(Z_i)_{i\geq 1}$ are i.i.d. copies of a random element Z_0 . Let $\bar{h}_n(Z_0) = \mathbb{E}[h_n(Z_0, Z_{1:n(1-1/k)}) \mid Z_0]$ and $\sigma_n^2 = \operatorname{Var}(\bar{h}_n(Z_0))$. If the following conditions hold:

•
$$\gamma_{loss}(h_n) = o(\sigma_n^2/n)$$
,

• the sequence of $(\bar{h}_n(Z_0) - \mathbb{E}[\bar{h}_n(Z_0)])^2 / \sigma_n^2$ is uniformly integrable,

then

$$\frac{\sqrt{n}}{\sigma_n} \left(\hat{R}_n - R_n \right) \stackrel{d}{\to} \mathcal{N}(0, 1).$$

Improvement upon previous literature

Dudoit and van der Laan, Theorem 3 (2005)

- Assumes a bounded loss function
- Excludes leave-one-out CV
- Requires the prediction rule to be loss-consistent for a risk-minimizing prediction rule

Improvement upon previous literature

Dudoit and van der Laan, Theorem 3 (2005)

- Assumes a bounded loss function
- Excludes leave-one-out CV
- Requires the prediction rule to be loss-consistent for a risk-minimizing prediction rule

Austern and Zhou, Theorem 1 (2020)

- Assumes variance parameter converging to a non-zero limit
- Requires o(1/n) mean-square stability
- Requires $o(1/n^2)$ second-order mean-square stability
- Assumes learning algorithms symmetric in the training points

k-fold test error confidence intervals

Goal

Construct an asymptotically-exact $(1-\alpha)\text{-confidence}$ interval for the unknown k-fold test error R_n

Goal

Construct an asymptotically-exact $(1-\alpha)\text{-confidence}$ interval for the unknown k-fold test error R_n

Confidence interval

Consider $\hat{\sigma}_n^2$ a variance estimator satisfying relative error consistency, $\hat{\sigma}_n^2/\sigma_n^2 \xrightarrow{p} 1$. With the CLT, $C_{\alpha} \triangleq \hat{R}_n \pm q_{1-\alpha/2} \hat{\sigma}_n/\sqrt{n}$

satisfies

$$\lim_{n \to \infty} \mathbb{P}(R_n \in C_\alpha) = 1 - \alpha$$

where $q_{1-\alpha/2}$ is the $(1-\alpha/2)\mbox{-quantile}$ of a standard normal distribution

Testing for algorithm improvement

Goal

Given a dataset $Z_{1:n}$, a k-fold partition $\{B'_j\}_{j=1}^k$, and two algorithms \mathcal{A}_1 , \mathcal{A}_2 for fitting prediction rules, test whether \mathcal{A}_1 improves upon \mathcal{A}_2 on the fold partition

Testing for algorithm improvement

Goal

Given a dataset $Z_{1:n}$, a k-fold partition $\{B'_j\}_{j=1}^k$, and two algorithms \mathcal{A}_1 , \mathcal{A}_2 for fitting prediction rules, test whether \mathcal{A}_1 improves upon \mathcal{A}_2 on the fold partition

Test

Define

$$h_n(Z_0, Z_B) = \ell(Y_0, \hat{f}_1(X_0; Z_B)) - \ell(Y_0, \hat{f}_2(X_0; Z_B)).$$

Consider $\hat{\sigma}_n^2$ a variance estimator satisfying relative error consistency, $\hat{\sigma}_n^2/\sigma_n^2 \xrightarrow{p} 1$. Test the null $H_0: R_n \ge 0$ against the alternative hypothesis $H_1: R_n < 0$. Asymptotically-exact level- α test is given by REJECT $H_0 \Leftrightarrow \hat{R}_n < q_\alpha \hat{\sigma}_n / \sqrt{n}$

where q_{α} is the $\alpha\text{-quantile}$ of a standard normal distribution

Consistent variance estimation

Want to find an estimator $\hat{\sigma}_n^2$ such that $\hat{\sigma}_n^2/\sigma_n^2 \xrightarrow{p} 1$ under weak conditions.

Definition (Within-fold variance estimator)

 $\hat{\sigma}_{n,in}^2$ is the average of the k within-fold empirical variances

Definition (All-pairs variance estimator)

$$\hat{\sigma}_{n,out}^2 \triangleq \frac{1}{n} \sum_{j=1}^k \sum_{i \in B'_j} (h_n(Z_i, Z_{B_j}) - \hat{R}_n)^2$$

Advantage: can also be used for leave-one-out cross-validation

Low computational cost

$$\hat{\sigma}_{n,in}^2$$
 and $\hat{\sigma}_{n,out}^2$ can be computed in $O(n)$ time

Consistent variance estimation

Theorem

Suppose $(Z_i)_{i\geq 1}$ are i.i.d. copies of a random element Z_0 . Let $\bar{h}_n(Z_0) = \mathbb{E}[h_n(Z_0, Z_{1:n(1-1/k)}) \mid Z_0]$ and $\sigma_n^2 = \operatorname{Var}(\bar{h}_n(Z_0))$. If the following conditions hold:

1)
$$\gamma_{loss}(h_n) = o(\sigma_n^2/n)$$
,

 ${\ensuremath{ \odot } }$ the sequence of $(\bar{h}_n(Z_0)-\mathbb{E}[\bar{h}_n(Z_0)])^2/\sigma_n^2$ is uniformly integrable,

then

$$\hat{\sigma}_{n,in}^2 / \sigma_n^2 \xrightarrow{L^1} 1.$$

Consistent variance estimation

Theorem

Suppose $(Z_i)_{i\geq 1}$ are i.i.d. copies of a random element Z_0 . Let $\bar{h}_n(Z_0) = \mathbb{E}[h_n(Z_0, Z_{1:n(1-1/k)}) \mid Z_0]$ and $\sigma_n^2 = \operatorname{Var}(\bar{h}_n(Z_0))$. If the following conditions hold:

1)
$$\gamma_{loss}(h_n) = o(\sigma_n^2/n)$$
,

 2 the sequence of $(\bar{h}_n(Z_0)-\mathbb{E}[\bar{h}_n(Z_0)])^2/\sigma_n^2$ is uniformly integrable,

then

$$\hat{\sigma}_{n,in}^2 / \sigma_n^2 \xrightarrow{L^1} 1.$$

If additionally:

then

$$\hat{\sigma}_{n,out}^2 / \sigma_n^2 \xrightarrow{L^1} 1.$$

Experiments – k-fold test error confidence intervals

$$C_lpha=\hat{R}_n\pm q_{1-lpha/2}\,\hat{\sigma}_n/\sqrt{n}$$
 with $lpha=0.05$





Figure: Test error coverage (top) and width (bottom) of 95% confidence intervals. Left: ℓ^2 -regularized logistic regression classifier. Right: Random forest regression.

Experiments – Testing for algorithm improvement

REJECT
$$H_0 \Leftrightarrow \hat{R}_n < q_\alpha \hat{\sigma}_n / \sqrt{n}$$
 with $\alpha = 0.05$

Our CV CLT procedure: valid size, most powerful



Figure: Size when testing H_1 : $\operatorname{Err}(\mathcal{A}_1) < \operatorname{Err}(\mathcal{A}_2)$ (top) and power when testing H_1 : $\operatorname{Err}(\mathcal{A}_2) < \operatorname{Err}(\mathcal{A}_1)$ (bottom) of level-0.05 tests for improved test error. Left: $\mathcal{A}_1 = \ell^2$ -regularized logistic regression, \mathcal{A}_2 = neural network classification. Right: \mathcal{A}_1 = random forest, \mathcal{A}_2 = ridge regression.





Thank you!

Cross-validation Confidence Intervals for Test Error Paper: https://arxiv.org/abs/2007.12671 Code: https://github.com/alexandre-bayle/cvci

Also in the paper:

- additional theoretical results
- experiments in the leave-one-out setting
- experiments illustrating the importance of stability